### Fitting Flexible Meta-Analytic Models with Structural Equation Modeling

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#### Overview of main functions in metaSEM package

- The metaSEM<sup>1</sup> package uses the OpenMx<sup>2</sup> and lavaan<sup>3</sup> packages to conduct meta-analysis, and the semPlot package<sup>4</sup> for visualization.
- Calculating effect sizes and their sampling covariance matrices
  - asyCov(), calEffSizes(), smdMES(), and smdMTS().
- Meta-analytic structural equation modeling (MASEM)
  - Two-stage SEM (with Wai Chan): tssem1(), tssem2(), and wls().
  - One-stage MASEM (with Suzanne Jak): osmasem() for correlation matrices and osmasem2() for covariance/correlation matrices and means (new!).
- SEM-based meta-analysis
  - Univariate and multivariate meta-analyses: meta() and metaFIML().
  - Three-level meta-analysis: meta3L() and meta3LFIML().
  - Fitting flexible models: sem() (today's talk).

<sup>&</sup>lt;sup>1</sup>Cheung, M. W.-L. (2015). metaSEM: An R package for meta-analysis using structural equation modeling. *Frontiers in Psychology*, 5(1521). https://doi.org/10.3389/fpsyg.2014.01521

<sup>&</sup>lt;sup>2</sup>Boker, S., Neale, M., Maes, H., Wilde, M., Spiegel, M., Brick, T., Spies, J., Estabrook, R., Kenny, S., Bates, T., Mehta, P., & Fox, J. (2011). OpenMx: An open source extended structural equation modeling framework. *Psychometrika*, 76(2), 306–317. https://doi.org/10.1007/s11336-010-9200-6

<sup>&</sup>lt;sup>3</sup>Rosseel, Y. (2012). Iavaan: An R package for structural equation modeling. Journal of Statistical Software, 48(2), 1–36.

<sup>&</sup>lt;sup>4</sup>Epskamp, S. (2015). semPlot: Unified Visualizations of Structural Equation Models. *Structural Equation Modeling: A Multidisciplinary Journal*, *22*(3), 474–483

#### Introduction

- Researchers usually learn meta-analytic models by comprehending the mathematics and algorithms.
- They often need to wait for programs, such as the metafor<sup>5</sup> package in R, to implement these models.
- Customizing meta-analytic models can be challenging, for example, fitting regression or mediation models with effect sizes.
- This talk introduces an SEM approach to address these concerns.

<sup>&</sup>lt;sup>5</sup>Viechtbauer, W. (2010). Conducting meta-analyses in R with the metafor package. Journal of Statistical Software, 36(3), 1–48. https://doi.org/10.18637/jss.v036.i03

#### A random-effects model with additive heterogeneity

- $y_i = \mu + u_i + e_i$ , with  $E(y_i) = \mu$ ,  $Var(e_i) = v_i$ , and  $Var(u_i) = \tau^2$ .
- A fixed- (or common-) effect model is a special case when  $\tau^2 = 0$ .
- SEM notations:

- Mean structure:  $\mu_i(\theta) = \mu$ .
- Variance structure:  $\Sigma_i(\theta) = \tau^2 + v_i.$



- □: an observed variable
- O: a latent variable
- $\triangle$ : a mean or intercept
- →: prediction or "cause," e.g., a regression coefficient
- ↔: association, e.g., a covariance or variance
- data.vi: assigning v<sub>i</sub> to the parameter (only in OpenMx)

#### A model with multiplicative heterogeneity

- The random-effects model with additive error is not the only model in the literature.
- There are debates whether the heterogeneity should be additive or multiplicative.<sup>6</sup>, <sup>7</sup>
- $y_i = \mu + \sqrt{\phi} e_i$ , with  $E(y_i) = \mu$  and  $Var(e_i) = v_i$ .
- $\hfill \phi$  is the scaling factor of the multiplicative heterogeneity variance.



<sup>&</sup>lt;sup>6</sup> Mawdsley, D., Higgins, J. P. T., Sutton, A. J., & Abrams, K. R. (2017). Accounting for heterogeneity in meta-analysis using a multiplicative model—An empirical study. *Research Synthesis Methods*, 8(1), 43–52. https://doi.org/10.1002/jrsm.1216

<sup>&</sup>lt;sup>7</sup>Stanley, T. D., & Doucouliagos, H. (2015). Neither fixed nor random: Weighted least squares meta-analysis. Statistics in Medicine, 34(13), 2116–2127. https://doi.org/10.1002/sim.6481

#### Additive or multiplicative or both?

 A hybrid model including both addictive and multiplicative heterogeneity has been proposed.<sup>8</sup>, <sup>9</sup>



 Mean structure: μ<sub>i</sub>(θ) = μ.
 Variance structure: Σ<sub>i</sub>(θ) = τ<sup>2</sup> + φv<sub>i</sub>.

<sup>&</sup>lt;sup>8</sup> Baker, R. D., & Jackson, D. (2013). Meta-analysis inside and outside particle physics: Two traditions that should converge? Research Synthesis Methods, 4(2), 109–124. https://doi.org/10.1002/jrsm.1065

<sup>&</sup>lt;sup>9</sup>Schmid, C. H. (2017). Heterogeneity: Multiplicative, additive or both? Research Synthesis Methods, 8(1), 119–120. https://doi.org/10.1002/jrsm.1223

#### Comparison of the above models

- Some of them are nested.
- AIC and BIC have been utilized for model comparison,<sup>10</sup> but it is unclear whether this is the best approach.



<sup>10</sup>Stanley, T. D., Ioannidis, J. P. A., Maier, M., Doucouliagos, H., Otte, W. M., & Bartoš, F. (2023). Unrestricted weighted least squares represent medical research better than random effects in 67,308 Cochrane meta-analyses. *Journal of Clinical Epidemiology*, 157, 53–58.

#### **Typical workflow**



## Example 1: A hybrid model with both additive and multiplicate heterogeneity<sup>11</sup>

- Effect size (correlation between organizational commitment and salesperson job performance): yi.
- Sampling variance: vi.
- Covariate (Individualism scores of the studies): xi

<sup>&</sup>lt;sup>11</sup> Jaramillo, F., Mulki, J. P., & Marshall, G. W. (2005). A meta-analysis of the relationship between organizational commitment and salesperson job performance: 25 years of research. *Journal of Business Research*, 58(6), 705–714. https://doi.org/10.1016/j.jbusres.2003.10.004

#### Step 0: Load the libraries and prepare the data

##		yi	vi	xi
##	1	0.02	0.005582124	-33.836066
##	2	0.12	0.004187101	9.163934
##	3	0.09	0.001756903	9.163934
##	4	0.20	0.005091713	-14.836066
##	5	0.08	0.006328468	9.163934
##	6	0.04	0.005537792	9.163934

### Step 1: Specify the model in R

 Readers may refer to the lavaan website for the lavaan syntax to specify structural equation models.

```
## Specify a hybrid model and call it "m1"
m1 <- "vi ~ mu*1
                ## Mean(yi) = mu
      vi ~~ 0*vi ## Measurement error is fixed at 0
      ## Additive error
      ui =~ 1*vi ## vi = 1*ui
      ui ~~ tau2*ui ## Var(ui) = tau2
      ## Multiplicative error
      ei =~ phi_sqrt*yi ## yi = phi_sqrt*ei
      ei ~~ data.vi*ei ## Var(ei) = vi
      phi := phi_sqrt<sup>2</sup> ## Define phi as a function of
                         ## (phi sqrt)^2
н
```

#### Step 2: Plot the conceptual model (optional)



### Step 3: Symbolically derive the model implied mean and variance structures (optional)

```
## Convert the model to a RAM specification
ram1 <- lavaan2RAM(m1, obs.variables = "yi", std.lv = FALSE)</pre>
## Get the model implied structures
impliedS(ram1)
## Model implied covariance matrix (Sigma):
## yi
## yi "data.vi*phi sqrt<sup>2</sup> + tau2"
## Model implied mean vector (Mu):
##
   yi
## 1 "muu"
```

#### Step 4: Fit the model with data

```
## Use the likelihood-based CI for phi=phi_sqrt^2
hybrid <- sem("Hybrid", RAM=ram1, data=dat, intervals.type="LB")
summary(hybrid)</pre>
```

```
## 95% confidence intervals: Likelihood-based statistic
## Coefficients:
       Estimate Std.Error lbound ubound z value Pr(>|z|)
##
## mu 0.1873704 NA 0.1488124 0.2257397
                                                     NA
                                                             NA
## phi_sqrt 1.2870404 NA -2.0902766 2.0872970 NA
                                                             NΑ
## tau2 0.0136638 NA 0.0039765 0.0304225 NA
                                                             NΑ
##
## Mxalgebras:
##
           lbound estimate ubound
## phi 1.805558e-44 1.656473 4.361042
##
## Information Criteria:
       df Penalty Parameters Penalty Sample-Size Adjusted
##
## ATC: -171.7740 -49.77403 -49.35298
## BIC: -294.2047 -43.44141
                                        -52.87879
##
## Number of subjects (or studies): 61
## Number of observed statistics: 61
## Number of estimated parameters: 3
## Degrees of freedom: 58
## -2 log likelihood: -55.77403
## OpenMx status1: 0 ("0" or "1": The optimization is considered fine.
## Other values may indicate problems.)
```

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# Step 5: Plot the model with the parameter estimates (optional)



#### Let's compare all four models

The hybrid model has the lowest AIC (best model), while the random-effects model's AIC is also very similar.

The fixed-effect model has the largest AIC (worst model).

anova(hybrid, random, multi, fixed)

##		base	comparison	ep	minus2LL	df	AIC	diffLL	
##	1	Hybrid	<na></na>	3	-55.77403	58	-49.77403	NA	
##	2	Hybrid	Random	2	-53.28211	59	-49.28211	2.491919	
##	3	Hybrid	Multiplicative	2	-44.62625	59	-40.62625	11.147781	
##	4	Hybrid	Fixed	1	129.06973	60	131.06973	184.843761	
##			р						
##	1		NA						
##	2	1.14432	21e-01						
##	3	3 8.413227e-04							
##	4	7.27255	59e-41						

## A mixed-effects model with a predictor $x_i$ as a design matrix

In a mixed-effects meta-analysis or meta-regression, the predictors are treated as a design matrix without any distribution assumption. This is known as the fixed-x approach.

$$y_i = \beta_0 + \beta_1 x_i + u_i + e_i.$$



### A mixed-effects model with a predictor $x_i$ as a variable

- In SEM, the predictors are usually treated as variables. This is known as the random-x approach.
- Pros: handling missing predictors with FIML estimation under the assumption of missing at random (MAR).
- Cons: predictors are assumed multivariate normal.

• 
$$y_i = \beta_0 + \beta_1 x_i + u_i + e_i$$
, with  $E(x_i) = \mu_x$ , and  $Var(x_i) = \sigma_x^2$ .

Mean structure:





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#### Using true effect size as a predictor (1)

- Effect sizes may be used as predictors rather than as outcome variables:
  - An effect size predicts another variable, e.g.,  $z_{citations} = \beta_0 + \beta_1 y_i + e_i^{12}$
  - An effect size predicts another effect size, e.g.,  $y_{\text{treatment benefit}} = \beta_0 + \beta_1 y_{\text{baseline risk}} + e_i^{13}$  and  $y_{\text{debunking effect}} = \beta_0 + \beta_1 y_{\text{misinformation effect}} + e_i$  in science-relevant misinformation<sup>14</sup>
- The estimated regression coefficients β<sub>1</sub>s are biased towards zero if we use the observed effect sizes as predictors.

<sup>&</sup>lt;sup>12</sup> Lortie, C. J., Aarssen, L. W., Budden, A. E., & Leimu, R. (2013). Do citations and impact factors relate to the real numbers in publications? A case study of citation rates, impact, and effect sizes in ecology and evolutionary biology. *Scientometrics*, 94(2), 675–682. https://doi.org/10.1007/s11192-012-0822-6

<sup>&</sup>lt;sup>13</sup>Arends, L. R., Hoes, A. W., Lubsen, J., Grobbee, D. E., & Stijnen, T. (2000). Baseline risk as predictor of treatment benefit: Three clinical meta-re-analyses. *Statistics in Medicine*, 19(24), 3497–3518. https://doi.org/10.1002/1097-0258(20001230)19:24-3497:JAID-SIM830>3.0.CO;2-H

<sup>&</sup>lt;sup>14</sup>Chan, M. S., & Albarracín, D. (2023). A meta-analysis of correction effects in science-relevant misinformation. Nature Human Behaviour, 7(9), 1514–1525. https://doi.org/10.1038/s41562-023-01623-8

#### Using true effect size as a predictor (2)

f<sub>i</sub> is the "true" effect size without sampling error:

 y<sub>i</sub> = f<sub>i</sub> + e<sub>i</sub>, with E(f<sub>i</sub>) = μ<sub>y</sub> and Var(f<sub>i</sub>) = τ<sup>2</sup><sub>y</sub>.

 An outcome variable:

 z<sub>i</sub> = β<sub>0</sub> + β<sub>1</sub>f<sub>i</sub> + e<sub>zi</sub>, with Var(e<sub>zi</sub>) = σ<sup>2</sup><sub>ex</sub>.



#### Meta-analysis beyond the linear relationship

- It is typically assumed that there is a linear relationship between the effect size and covariates.
- Additionally, it is usually assumed that the heterogeneity variance (τ<sup>2</sup>) remains constant across different levels of the covariates.
- However, it is crucial to acknowledge that there are situations where one or both of these assumptions may not hold true.

#### An example of nonlinear mean structure<sup>15</sup>

#### Figure 2

Life Span Trends for Rank-Order Stability Estimates (r) for All Traits and the Big Five Separately



Note: The first panel plots results for the full data set, and the subsequent panels plot results for Extraversion. Agreeableness, Conscientiousness. Emotional Stability, and Openness, in that order: Effect sizes are plotted in addition to the best-fitting optime model and scaled relative to the weight the effect size carried the analysis, with larger plotting characters carrying more weight. Effect sizes represented as a circle are from previous meta-analyses, and effect sizes represented as a triangle are from the newly coded data. Shading around the trend line reflexts the 95% confidence interval. Gen. = general personality effect size; Eta . = Extraversion, Agre = Agreeadbeness; Chas = Goordiscinuossiss; Eta no. = Entointical Shability; Opn = Openness.

<sup>15</sup>Bleidorn, W., Schwaba, T., Zheng, A., Hopwood, C. J., Sosa, S. S., Roberts, B. W., & Briley, D. A. (2022). Personality stability and change: A meta-analysis of longitudinal studies. *Psychological Bulletin*, 148(7–8), 588–619. https://doi.org/10.1037/bul0000365

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### An example of non-constant heterogeneity variance<sup>16</sup>



Figure 2. Scatterplot displaying the relation between effect size (i.e., standardized mean change d per year,  $d_{ywd}$ ) and age (i.e., mean age of sample at the center of the observed time interval). The figure also shows the locally weighted smoothing (LOESS) curve across age.

<sup>16</sup>Orth, U., Erol, R. Y., & Luciano, E. C. (2018). Development of self-esteem from age 4 to 94 years: A meta-analysis of longitudinal studies. *Psychological Bulletin*, 144(10), 1045–1080. https://doi.org/10.1037/bul0000161

#### Example 2: Location-scale<sup>17</sup> and nonlinear models

- A location-scale model:
  - $\mathbf{y}_i = \underbrace{\mu_i}_{i} + \underbrace{u_i + e_i}_{i} \quad .$
  - Mean structure Variance structure Mean structure:  $\mu_i(\theta) = \mu_i = \beta_0 + \beta_1 x_i$ .
  - Variance structure:  $\Sigma_i(\theta) = \operatorname{Var}(u_i + e_i) = \exp(\alpha_0 + \alpha_1 x_i) + v_i$ , where  $\exp(\alpha_0)$  is the heterogeneity variance when  $x_i = 0$ .
- Notes:
  - The predictors in the mean and variance structures can be different.
  - A nonlinear structure may also be applied on the mean structure.

<sup>&</sup>lt;sup>17</sup> Viechtbauer, W., & López-López, J. A. (2022). Location-scale models for meta-analysis. Research Synthesis Methods, 13(6), 697–715. https://doi.org/10.1002/jrsm.1562

### Step 1: Specify a location-scale model with 6 lines of R code

### Step 2: Symbolically derive the model implied structures

## Convert the model to a RAM specification
ram2 <- lavaan2RAM(m2, obs.variables="yi", std.lv=FALSE)</pre>

## Get the model implied structures
## We need to replace the constraints with the new parameters
impliedS(ram2, replace.constraints=TRUE)

## Model implied covariance matrix (Sigma):

```
## yi
## yi "data.vi + exp(a0 + a1*data.xi)"
## Model implied mean vector (Mu):
## yi
## 1 "b0 + b1*data.xi"
```

#### Step 3: Fit the model with data

## We need to replace the constraints with the new parameters
fit2 <- sem(RAM=ram2, data=dat, replace.constraints=TRUE)
summary(fit2)</pre>

```
## 95% confidence intervals: z statistic approximation (robust=FALSE)
## Coefficients:
##
        Estimate Std.Error lbound ubound z value Pr(>|z|)
## a0 -4.11206737 0.24483731 -4.59193967 -3.63219506 -16.7951 <2e-16 ***
## a1 0.00541492 0.01314676 -0.02035226 0.03118209 0.4119 0.6804
## b0 0.18596386 0.01908419 0.14855954 0.22336818 9.7444 <2e-16 ***
## b1 -0.00129620 0.00090683 -0.00307355 0.00048115 -1.4294 0.1529
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Information Criteria:
##
       df Penalty Parameters Penalty Sample-Size Adjusted
## AIC: -171.4105 -49.41046 -48.69617
## BIC: -291.7303 -40.96696
                                          -53.55013
##
## Number of subjects (or studies): 61
## Number of observed statistics: 61
## Number of estimated parameters: 4
## Degrees of freedom: 57
## -2 log likelihood: -57.41046
## OpenMx status1: 0 ("0" or "1": The optimization is considered fine.
## Other values may indicate problems.)
```

#### A random-effects bivariate meta-analysis

The previous univariate models can be extended to bivariate models.



### A random-effects bivariate meta-analysis with common between and within correlation

- One difficult aspect of applying multivariate meta-analysis is that information for calculating the sampling correlation may be missing.
- Riley et al.<sup>18</sup> proposed a model that assumes the between- and within-study correlations are the same.



<sup>18</sup> Riley, R. D., Thompson, J. R., & Abrams, K. R. (2008). An alternative model for bivariate random-effects meta-analysis when the within-study correlations are unknown. *Biostatistics*, 9(1), 172–186. https://doi.org/10.1093/biostatistics/kxm023

## Strengths and limitations of the SEM-based meta-analysis

#### Strengths:

- Learning meta-analysis using graphical models
- Deriving model-implied means and covariances
- Fitting proposed models to data
- Imposing linear and nonlinear constraints on the parameters
- Creating functions of parameters and their confidence intervals, such as  $R^2$  and indirect effect a \* b
- Handling missing data with the full information maximum likelihood (FIML) estimation method
- Extending to multivariate meta-analysis and even MASEM!

Limitations:

- Fitting the data with FIML only
- Using a z test rather than a t test (but easy to fix)
- Having to implement the three-level meta-analysis and robust variance estimation (RVE).

#### Conclusion

- This approach can be used to fit novel meta-analytic models, which are not available in the existing meta-analysis software yet.
- The preprint is available at https://osf.io/preprints/psyarxiv/w9pc6.
- The R code for the analyses is available at https://github.com/mikewlcheung/code-inarticles/tree/master/Cheung%202024.
- Questions and comments are welcome!